

Fastest mixing Markov chain on a path ^{*}

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Abstract

We consider the problem of assigning transition probabilities to the edges of a path, so the resulting Markov chain or random walk mixes as rapidly as possible. In this note we prove that fastest mixing is obtained when each edge has a transition probability of $1/2$. Although this result is intuitive (it was conjectured in [7]), and can be found numerically using convex optimization methods [2], we give a self-contained proof.

In [2], the authors consider the problem of assigning transition probabilities to the edges of a connected graph in such a way that the associated Markov chain mixes as rapidly as possible. We show that this problem can be solved, at least numerically, using tools of convex optimization, in particular, semidefinite programming [9, 3]. The present note presents a simple, self contained example where the optimal Markov chain can be identified analytically.

Consider a path with $n \geq 2$ nodes, labeled $1, 2, \dots, n$, with $n - 1$ edges connecting pairs of adjacent nodes, and a loop at each node, as shown in figure 1. We consider a Markov chain (or random walk) on this path, with transition probability from node i to node j denoted P_{ij} . The requirement that transitions can only occur along an edge or loop of the path is equivalent to $P_{ij} = 0$ for $|i - j| > 1$, *i.e.*, P is a tridiagonal matrix. Since P_{ij} are transition probabilities, we have $P_{ij} \geq 0$, and $\sum_j P_{ij} = 1$, *i.e.*, P is a stochastic matrix. This can be expressed as $P\mathbf{1} = \mathbf{1}$, where $\mathbf{1}$ is the vector with all components one.

We will consider symmetric transition probabilities, *i.e.*, those that satisfy $P_{ij} = P_{ji}$. Thus, P is a symmetric, (doubly) stochastic, tridiagonal matrix. Since $P\mathbf{1} = \mathbf{1}$, we have $(\mathbf{1}/n)^T P = \mathbf{1}/n$, which means that the uniform distribution, given by $\mathbf{1}^T/n$, is stationary.

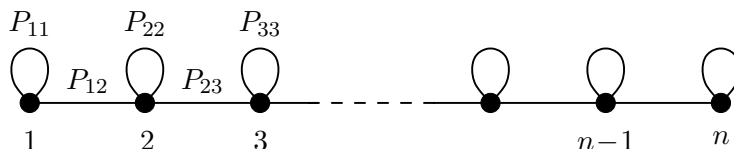


Figure 1: A path with loops at each node, with transition probabilities labeled.

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represents a direct network connection between the adjacent processors. Processor i has a job queue or load $q_i(t)$ (which we approximate as a positive real number) at time t . The goal is to shift jobs across the links, at each step, in such a way as to balance the load. In other words, we would like to have $q_i(t) \rightarrow \bar{q}$ as $t \rightarrow \infty$, where $\bar{q} = (1/n) \sum_i q_i(0)$ is the average of the initial queues. We ignore the reduction in the queues due to processing (or equivalently, assume that the load balancing is done before the processing begins). We use the following simple scheme to balance the load: at each step, we compute the load imbalance, $q_{i+1}(t) - q_i(t)$, across each link. We then transfer a fraction $\theta_i \in [0, 1]$ of the load imbalance from the more loaded to the less loaded processor. We must have $\theta_i + \theta_{i+1} \leq 1$, to ensure that we are not asked to transfer more than the load on a processor to its neighbors. It can be shown that if θ_i are positive, and satisfy $\theta_i + \theta_{i+1} \leq 1$, then this iterative scheme achieves asymptotic balanced loads, *i.e.*, $q_i(t) \rightarrow \bar{q}$ as $t \rightarrow \infty$. The problem is to find the fractions θ_i that result in the fastest possible load balancing.

It turns out that this optimal iterative load balancing problem is identical to the problem of finding the fastest mixing Markov chain on a path, with $P_{i,i+1} = \theta_i$. In particular, the evolution of the loads at the processors is given by $q(t) = P^t q(0)$. The speed of convergence of $q(t)$ to $\bar{q}\mathbf{1}$ is given by the second-largest eigenvalue modulus $\mu(P)$. By the basic result in this paper, the fastest possible load balancing is accomplished by shifting one-half of the load imbalance on each edge from the more loaded to the less loaded processor. More discussion of this load balancing problem can be found in [7].

We now proceed to prove the basic result.

Lemma. *Let $P \in \mathbf{R}^{n \times n}$ be a symmetric stochastic matrix. Then we have*

$$\mu(P) = \|P - (1/n)\mathbf{1}\mathbf{1}^T\|_2,$$

where $\|\cdot\|_2$ denotes the spectral norm (maximum singular value).

Proof. To see this, we note that $\mathbf{1}$ is the eigenvector of P associated with the eigenvalue $\lambda_1 = 1$. Therefore the eigenvalues of $P - \mathbf{1}\mathbf{1}^T/n$ are $0, \lambda_2, \dots, \lambda_n$. Since $P - \mathbf{1}\mathbf{1}^T/n$ is symmetric, its spectral norm is equal to the maximum magnitude of its eigenvalues, *i.e.*, $\max\{\lambda_2, -\lambda_n\}$, which is $\mu(P)$. \square

Lemma. *Let $P \in \mathbf{R}^{n \times n}$ be a symmetric stochastic matrix, and suppose $y, z \in \mathbf{R}^n$ satisfy*

$$\mathbf{1}^T y = 0, \quad \|y\|_2 = 1, \tag{2}$$

$$(z_i + z_j)/2 \leq y_i y_j \text{ for } P_{ij} \neq 0. \tag{3}$$

Then we have $\mu(P) \geq \mathbf{1}^T z$.

Proof. For any P, y and z that satisfy the assumptions in the lemma, we have

$$\begin{aligned} \mu(P) &= \|P - (1/n)\mathbf{1}\mathbf{1}^T\|_2 \\ &\geq y^T (P - (1/n)\mathbf{1}\mathbf{1}^T) y \\ &= y^T P y \\ &= \sum_{i,j} P_{ij} y_i y_j \end{aligned}$$

$$\begin{aligned}
&\geq \sum_{i,j} (1/2)(z_i + z_j)P_{ij} \\
&= (1/2)(z^T P \mathbf{1} + \mathbf{1}^T P z) \\
&= \mathbf{1}^T z.
\end{aligned}$$

The first inequality follows from the assumption $\|y\|_2 = 1$ and the first lemma. The second inequality follows from the assumption (3), and $P_{ij} \geq 0$. \square

Theorem. *The matrix P^* , given in (1), attains the smallest value of μ , $\cos(\pi/n)$, among all symmetric stochastic tridiagonal matrices.*

Proof. The result is clear for $n = 2$. We assume now that $n > 2$. The eigenvalues and associated orthonormal eigenvectors of P^* are

$$\begin{aligned}
\lambda_1 &= 1, & v_0 &= (1/\sqrt{n})\mathbf{1} \\
\lambda_j &= \cos\left(\frac{(j-1)\pi}{n}\right), & v_j(k) &= \sqrt{\frac{2}{n}} \cos\left(\frac{(2k-1)(j-1)\pi}{2n}\right), & j &= 2, \dots, n \\
& & & & k &= 1, \dots, n.
\end{aligned}$$

(See, *e.g.*, [8, §16.3].) Therefore we have

$$\mu(P^*) = \lambda_2 = -\lambda_n = \cos(\pi/n).$$

We show that this is the smallest μ possible by constructing a pair y and z that satisfy the assumptions (2) and (3) in the second lemma, for any symmetric tridiagonal stochastic matrix P , with $\mathbf{1}^T z = \cos(\pi/n)$.

We take $y = v_2$, so the assumptions (2) in the second lemma clearly hold. We take z to be

$$z_i = \frac{1}{n} \left[\cos\left(\frac{\pi}{n}\right) + \cos\left(\frac{(2i-1)\pi}{n}\right) \right] / \cos\left(\frac{\pi}{n}\right), \quad i = 1, \dots, n.$$

It is easy to verify that $\mathbf{1}^T z = \cos(\pi/n)$.

It remains to check that y and z satisfy (3) for any symmetric tridiagonal matrix P . Let's first check the superdiagonal entries. For $i = 1, \dots, n-1$, we have

$$\begin{aligned}
\frac{z_i + z_{i+1}}{2} &= \frac{1}{n} \left[\cos\left(\frac{\pi}{n}\right) + \frac{1}{2} \left(\cos\left(\frac{(2i-1)\pi}{n}\right) + \cos\left(\frac{(2i+1)\pi}{n}\right) \right) \right] / \cos\left(\frac{\pi}{n}\right) \\
&= \frac{1}{n} \left[\cos\left(\frac{\pi}{n}\right) + \cos\left(\frac{2i\pi}{n}\right) \right] \\
&= \frac{2}{n} \cos\left(\frac{(2i-1)\pi}{2n}\right) \cos\left(\frac{(2i+1)\pi}{2n}\right) = y_i y_{i+1}.
\end{aligned}$$

Therefore equality always holds for the superdiagonal (and subdiagonal) entries. For the diagonal entries, we need to check $(z_i + z_i)/2 = z_i \leq y_i^2$, *i.e.*,

$$\cos\left(\frac{\pi}{n}\right) + \cos\left(\frac{(2i-1)\pi}{n}\right) / \cos\left(\frac{\pi}{n}\right) \leq 2 \cos^2\left(\frac{(2i-1)\pi}{2n}\right) = 1 + \cos\left(\frac{(2i-1)\pi}{n}\right)$$

for $i = 1, \dots, n$, which is equivalent to

$$\left[1 - \cos\left(\frac{\pi}{n}\right)\right] \left[1 - \cos\left(\frac{(2i-1)\pi}{n}\right) / \cos\left(\frac{\pi}{n}\right)\right] \geq 0, \quad i = 1, \dots, n.$$

But this is certainly true because

$$\cos\left(\frac{(2i-1)\pi}{n}\right) \leq \cos\left(\frac{\pi}{n}\right), \quad i = 1, \dots, n.$$

This completes the proof. □

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